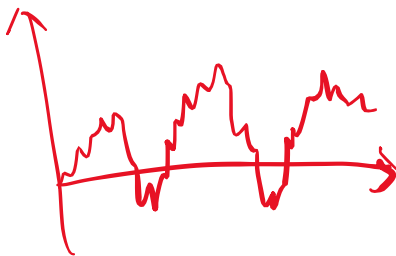
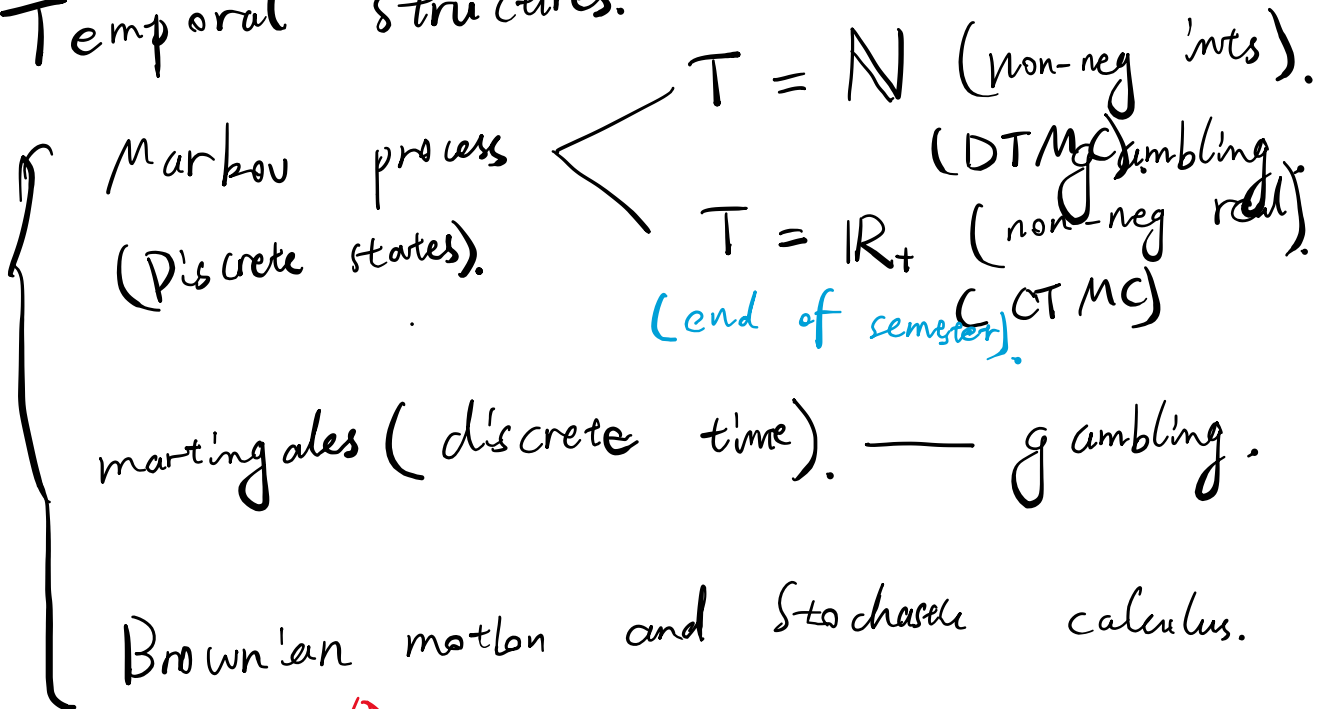


logistics.

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$(X_t : t \in T)$. T a period of time.

Temporal structures.



Def. Discrete-time, discrete-state, time-homogeneous MC.

(X_0, X_1, X_2, \dots) . $X_i \in S$.

(i) State space S
(finite or countably infinite).

(ii) Initial distribution $(\nu_i)_{i \in S}$

$$X_0 \sim \nu$$
$$(\mathbb{P}(X_0 = i) = \nu_i \text{ for } i \in S).$$

(iii) Transition probabilities.

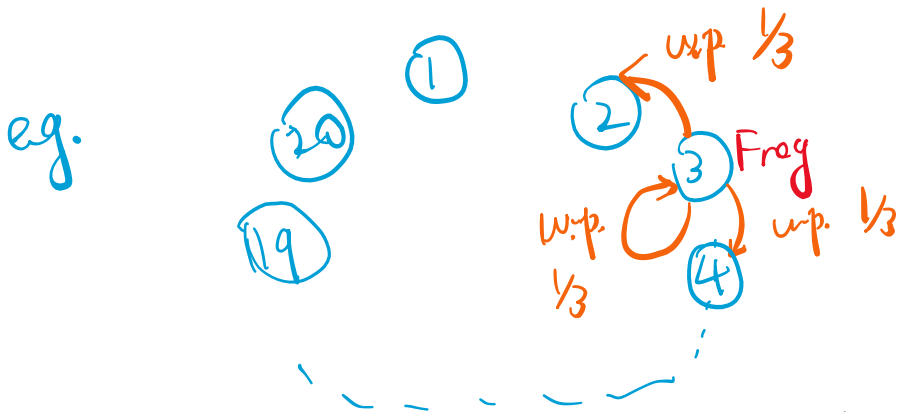
$$P = (P_{ij})_{i, j \in S}$$

Does not depend on time.

(In general, time inhomogeneous md.)

$$P_{ij} = \mathbb{P}(X_{t+1} = j \mid X_t = i)$$

(for $i, j \in S$).



$X_i :=$ index of the lily pad frog sitting on.

$$S = \{1, 2, 3, \dots, 20\}$$

$$V_i = \begin{cases} 1 & i=3, \\ 0 & i \neq 3. \end{cases}$$

$$P_{ij} = \begin{cases} \frac{1}{3} & \text{when } i=j \text{ or} \\ & i = (j+1) \bmod 20 \text{ or} \\ & i = (j-1) \bmod 20 \\ 0 & \text{otherwise.} \end{cases}$$



"transition matrix".

eg. Coin tossing.

At each time $t = 1, 2, 3, \dots$, toss a coin.

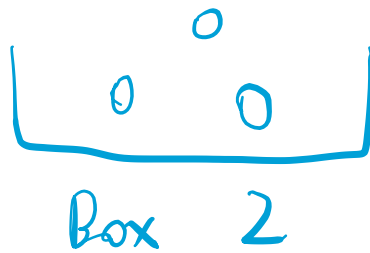
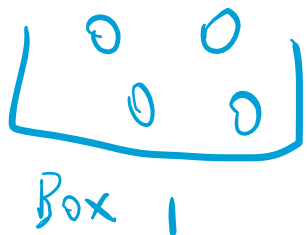
$X_t :=$ # heads in first t rounds.

$$S = \{0, 1, 2, \dots\}$$

$$v(0) = v_0 = 1, \quad v_i = 0 \text{ (for } i \geq 1\text{)}.$$

$$P_{ij} = \begin{cases} \frac{1}{2} & (i=j, \text{ or } i+1=j) \\ 0 & (\text{otherwise}). \end{cases}$$

eg. Ehrenfest's Urn.



d balls
in total.
 $d = a + b$

At each time.

— Randomly select a ball (uniformly).

— Move it to the opposite side.

Start w/ a balls in 1 and b balls in 2.

$X_t :=$ # of balls in box 1
at time t

$= \{0, 1, 2, \dots, d\}$

$= 1$, $v_i = 0$ for $(i \neq a)$

$P_{ij} = P(X_{t+1} = j \mid X_t = i)$

$X_{t+1} = \begin{cases} X_t + 1 & \text{w.p. } 1 - \frac{i}{d} \\ X_t - 1 & \text{w.p. } \frac{i}{d} \end{cases}$

$P_{ij} = \begin{cases} 1 - \frac{i}{d} & \text{when } j = i+1 \\ \frac{i}{d} & \text{when } j = i-1 \\ 0 & \text{otherwise.} \end{cases}$

Non-example if the move depends on history before time t .

Important property of MC
"Markov property".

$$\begin{aligned} & \mathbb{P}(X_t = j \mid X_0 = i_0, X_1 = i_1, X_2 = i_2, \dots, X_{t-1} = i_{t-1}) \\ &= \mathbb{P}(X_t = j \mid X_{t-1} = i_{t-1}) = P_{i_{t-1}j} \end{aligned}$$

Corollary.

$$\begin{aligned} & \mathbb{P}(X_0 = i_0, X_1 = i_1, \dots, X_n = i_n) \\ &= \mathbb{P}(X_0 = i_0) \cdot \mathbb{P}(X_1 = i_1 \mid X_0 = i_0) \cdot \mathbb{P}(X_2 = i_2 \mid X_0 = i_0, X_1 = i_1) \\ & \quad \dots \cdot \mathbb{P}(X_n = i_n \mid X_0 = i_0, \dots, X_{n-1} = i_{n-1}). \end{aligned}$$

$$= \nu_{i_0} \cdot P_{i_0 i_1} \cdot P_{i_1 i_2} \cdot \dots \cdot P_{i_{n-1} i_n}.$$

$$\mathbb{P}(X_0 = i_0, X_1 = i_1, X_2 = i_2) = \nu_{i_0} \cdot P_{i_0 i_1} \cdot P_{i_1 i_2}$$

$$\mathbb{P}(X_0 = i_0, X_2 = i_2) = \nu_{i_0} \cdot \sum_{i_1 \in S} P_{i_0 i_1} \cdot P_{i_1 i_2}$$

Matrix multiplication.

$$\mathbb{P}(X_2 = i_2 \mid X_0 = i_0) = [P^2]_{i_0, i_2}.$$

Detail A, B are infinite-dim matrices (countably inf S).

$$A = (a_{ij})_{i,j \in S} \quad B = (b_{ij})_{i,j \in S}$$

$$[A \cdot B]_{ij} = \sum_{k \in S} a_{ik} b_{kj}$$

infinite sum, convergence undem in general
but convergence is true
when A, B are prob. transition matrices.

In general, for integer $k \geq 0$.

$$\mathbb{P}(X_k = j \mid X_0 = i) = [P^k]_{ij}$$

$$\mathbb{P}(X_k = j) = \underbrace{[v \cdot P^k]}_{\text{(Row vector)}} j = \mathbb{P}(X_{n+k} = j \mid X_n = i)$$

Def. $P_{ij}^{(n)} := \mathbb{P}(X_n = j \mid X_0 = i)$

$$\left(= \mathbb{P}(X_{n+m} = j \mid X_m = i) \right)$$

for $i, j \in S$.

We have $P_{ij}^{(n)} = [P^n]_{ij}$.

$$P_{ij}^{(m+n)} = [P^{m+n}]_{ij}$$

$$= [P^m \cdot P^n]_{ij}$$

$$= \sum_{k \in S} P_{ik}^{(m)} \cdot P_{kj}^{(n)}$$

"Chapman-Kolmogorov Eq"

Recurrence and transience.

Def. $N(i) :=$ total number of times that MC visits state i

$$= \sum_{t=1}^{+\infty} \mathbb{1}\{X_t = i\}$$

(r.v., may be infinite in some cases)

$$f_{ij} := P(N(j) \geq 1 \mid X_0 = i)$$

(Prob. reach j when starting from i).

f_{ii} : prob of returning to i
if we start from i .

Def. A state i is $\begin{cases} \text{recurrent} & \text{if } f_{ii} = 1 \\ \text{transient} & \text{if } f_{ii} < 1 \end{cases}$

Fact. $\mathbb{P}_i(N(i) \geq k) = f_{ii}^k$

$= \mathbb{P}(N(i) \geq k \mid X_0 = i)$

(Not to be confused w/ $\mathbb{P}_x(X=x)$, not used in my lec.)

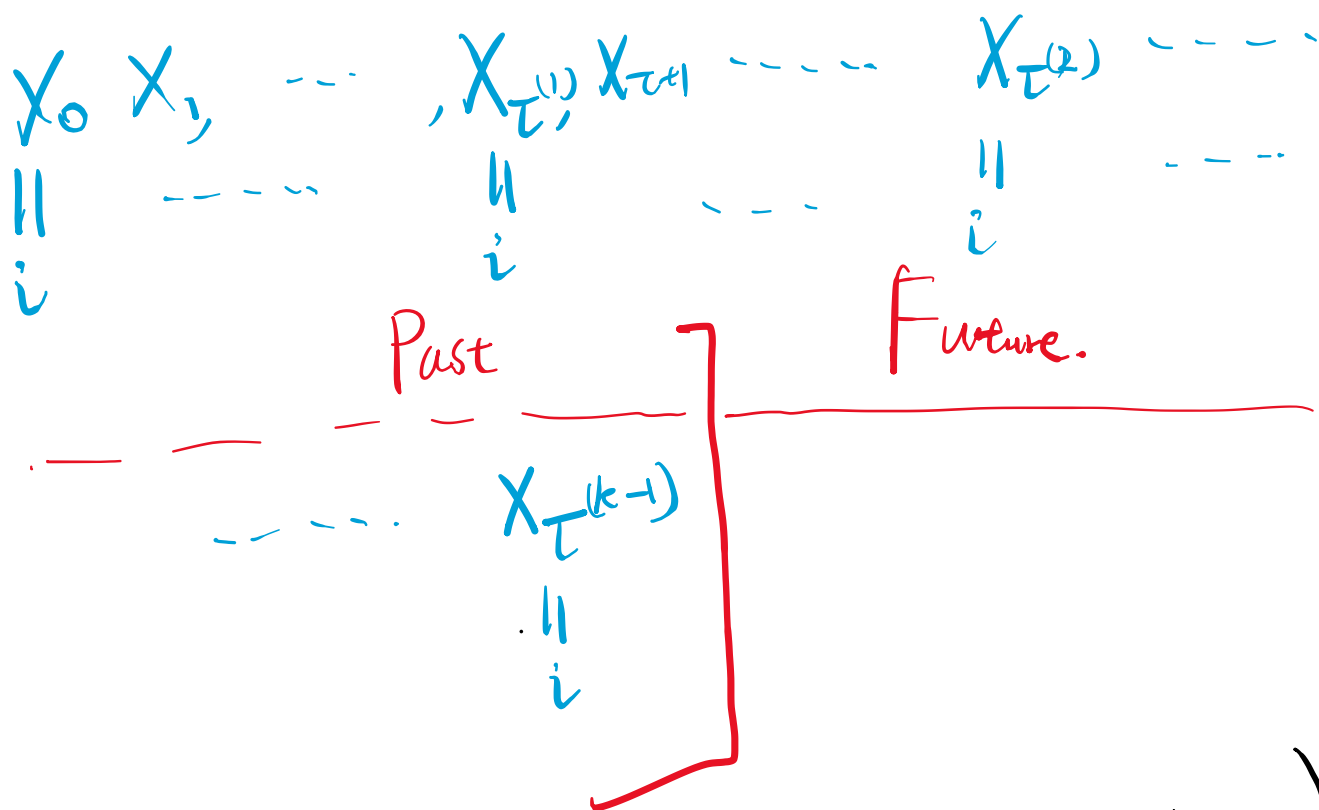
"Proof." Induction. $k=1$. By def.
Assuming the claim is true for $k-1$.

$$\mathbb{P}_i(N(i) \geq k)$$

$$= \mathbb{P}_i(N(i) \geq k \mid N(i) \geq k-1) \cdot \mathbb{P}_i(N(i) \geq k-1)$$

Need to show $= f_{ii}$.

$= f_{ii}^{k-1}$ by induction.



$$\begin{aligned}
 & \mathbb{P}(\text{visit } i \text{ again} \mid X_0, X_1, \dots, X_{\tau^{(k-1)}}) \\
 &= \mathbb{P}(\text{visit } i \text{ again} \mid X_{\tau^{(k-1)}} = i) \\
 &= \mathbb{P}(\text{visit } i \text{ again} \mid X_0 = i) = f^i.
 \end{aligned}$$

"Strong Markov property".

Let τ be the hitting time of state i ,

then we have

$$(X_\tau, X_{\tau+1}, \dots) \stackrel{d}{=} (X_0, X_1, \dots).$$

Also applies to

$\tau^{(k)}$ = k -th hitting time of i^* .

By strong Markov property,

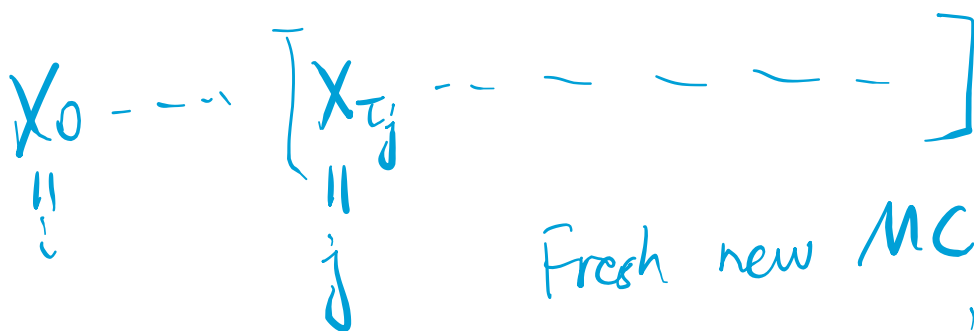
$$\mathbb{P}_i(N(i) \geq k \mid N(i) \geq k-1) = f_{ii}.$$

Def. hitting time of i

$$\tau_i := \inf \{ t \geq 1 : X_t = i \}$$

Corollary.

$$\mathbb{P}_i(N(j) \geq k) = f_{ij} \cdot f_{jj}^{k-1}.$$



Fresh new MC starting from j .

$$\begin{aligned}
& P_i(N(j) \geq k) \\
&= P_i(N(j) \geq 1) \cdot \underbrace{P_i(N(j) \geq k \mid N(j) \geq 1)} \\
&= f_{ij} \cdot P_j(N(j) \geq k-1) \\
&= f_{ij} \cdot f_{jj}^{k-1}
\end{aligned}$$

Notation

$$\begin{aligned}
& P_i(N(j) \geq k \mid N(j) \geq 1) \\
&= P(N(j) \geq k \mid N(j) \geq 1, X_0 = i). \\
&= P(N(j) \geq k \mid N(j) \geq 1) \\
&= P(N(j) \geq k-1 \mid X_0 = j).
\end{aligned}$$

Conclusion. (starting from i)

If $f_{ii} = 1$ then $N(i) = +\infty$
w.p. 1

If $f_{ii} < 1$ then

$$P(N(i) = k) = f_{ii}^k - f_{ii}^{k+1}$$

(Geometric distribution).

$$\text{e.g. } E_i[N(i)] = \begin{cases} \frac{f_{ij}}{1-f_{ij}} & (f_{ij} < 1) \\ +\infty & (f_{ij} = 1, f_{ij} > 0) \\ 0 & (f_{ij} = 0). \end{cases}$$

Question: can we determine
recurrence/transience by computing
the using $(P_{ij})_{i,j \in S}$.

Recurrence State Theorem.

State i is recurrent if and only if

$$\sum_{n=1}^{+\infty} P_{ii}^{(n)} = +\infty.$$

Proof.

$$\sum_{n=1}^{+\infty} P_{ii}^{(n)}$$

$$= \sum_{n=1}^{+\infty} P_i(X_n = i)$$

$$= \sum_{n=1}^{+\infty} \mathbb{E}_i[\mathbb{1}_{X_n=i}].$$

(Fubini - Tonelli)

$$= \mathbb{E}_i\left[\sum_{n=1}^{+\infty} \mathbb{1}_{X_n=i}\right].$$

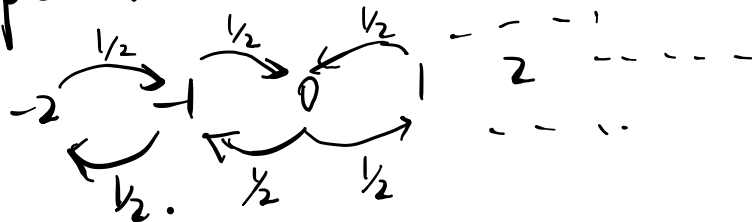
$$= \mathbb{E}_i[N(i)] = \begin{cases} \frac{f_{ii}}{1 - f_{ii}} & f_{ii} < 1 \\ +\infty & f_{ii} = 1. \end{cases}$$

$$f_{ii} < 1$$

$$f_{ii} = 1.$$

(By-product, able to compute f_{ii} from $(P_{ij})_{i,j \in S}$)

eg. Simple Random Walk.



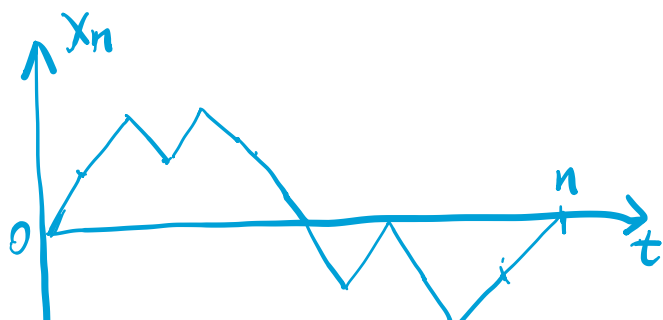
$$P_{i(i+1)} = P_{i(i-1)} = \frac{1}{2} \quad \text{for } i \in S = \mathbb{Z}.$$

Question: whether $f_{00} = 1$?

Suffices to check (by recurrent state thm)

$$\sum_{n=1}^{+\infty} P_{00}^{(n)}$$

$$P_{00}^{(n)} = 0 \quad (\text{if } n \text{ is odd}).$$



Total # paths = 2^n (for each step \uparrow or \downarrow).

paths that go back to 0 at time $n = \binom{n}{n/2} = \frac{n!}{(n/2)!^2}$

$(\frac{n}{2} \uparrow$'s and $\frac{n}{2} \downarrow$'s for n steps)

Stirling's approx $n! \approx \sqrt{2\pi n} \cdot (\frac{n}{e})^n$.

$$P_{00}^{(n)} \approx \sqrt{\frac{2}{\pi}} \cdot \frac{1}{\sqrt{n}}$$

$$\sum_{n=1}^{+\infty} P_{00}^{(n)} = +\infty$$

So 0 is recurrent.